

VALORE ATTESO e MATRICE DI COVARIANZA

$$(X, Y), \begin{pmatrix} X \\ Y \end{pmatrix}$$

$$\text{VALORE ATTESO: } (E[X], E[Y]) \\ E[(X, Y)]$$

FUNZIONE DI UN VETTORE ALEATORIO

$$h: \mathbb{R}^2 \rightarrow \mathbb{R}$$

$$Z = h(X, Y)$$

Teorema

(X, Y) vettore aleatorio discreto con densità discreta congiunta con $P(X, Y)$

$$E[Z] = E[h(X, Y)] = \sum_{i,j} h(x_i, y_j) P_{(X,Y)}(x_i, y_j)$$

COROLLARIO

(X, Y) vettore aleatorio discreto (cioè, X e Y sono v.a. discrete).

$$X \perp\!\!\!\perp Y \xRightarrow{\text{~~non~~}} \mathbb{E}[XY] = \mathbb{E}[X]\mathbb{E}[Y].$$

N.B.

$$X \perp\!\!\!\perp Y : \mathbb{P}(X \in A, Y \in B) = \mathbb{P}(X \in A) \mathbb{P}(Y \in B), \quad A, B \subset \mathbb{R}$$

$$X \perp\!\!\!\perp Y \iff P_{(X,Y)}(x_i, y_j) = P_X(x_i) P_Y(y_j), \quad \forall i, \forall j$$

then: $h(x, y) = xy$

$X \perp\!\!\!\perp Y$

$$\mathbb{E}[XY] \stackrel{\text{~~non~~}}{=} \sum_{i,j} x_i y_j P_{(X,Y)}(x_i, y_j) \stackrel{\text{~~non~~}}{=} \sum_{i,j} x_i y_j P_X(x_i) P_Y(y_j)$$

$$= \left(\sum_i x_i P_X(x_i) \right) \left(\sum_j y_j P_Y(y_j) \right) = \mathbb{E}[X] \mathbb{E}[Y].$$

OSS.

$$X \perp\!\!\!\perp Y \iff$$

$$f(X) \perp\!\!\!\perp g(Y)$$

Si può dimostrare che
 $X \perp\!\!\!\perp Y \iff$

$$\mathbb{E}[f(X)g(Y)] = \mathbb{E}[f(X)]\mathbb{E}[g(Y)]$$

MATRICE DI COVARIANZA

$$\mathbb{E}[(X, Y)] = (\mathbb{E}[X], \mathbb{E}[Y])$$

$(X, Y) :$

$$\begin{pmatrix} a_{11} & a_{12} \\ a_{21} & a_{22} \end{pmatrix} = \begin{pmatrix} \text{Var}(X) & \text{Cov}(X, Y) \\ \text{Cov}(Y, X) & \text{Var}(Y) \end{pmatrix}$$

$$(X, Y) = (X_1, X_2) \begin{pmatrix} \text{Cov}(X_1, X_1) & \text{Cov}(X_1, X_2) \\ \text{Cov}(X_2, X_1) & \text{Cov}(X_2, X_2) \end{pmatrix}$$

$$a_{ij} = \text{Cov}(X_i, X_j)$$

$$1) \text{Cov}(X_1, X_1) = \text{Var}(X_1) \quad | \quad 2) \text{Cov}(X, Y) = \text{Cov}(Y, X)$$

Definizione

X e Y sono v.a. discrete. La COVARIANZA di X e Y è data da

$$\begin{aligned}\text{Cov}(X, Y) &= \mathbb{E}[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])] \\ &= \sum_{i,j} (x_i - \mathbb{E}[X])(y_j - \mathbb{E}[Y]) P_{(X,Y)}(x_i, y_j)\end{aligned}$$

Se $\text{Cov}(X, Y) = 0$, le v.a. X e Y si dicono

SCORRELATE.

Se $\text{Var}(X) > 0$ e $\text{Var}(Y) > 0$, definiamo il COEFFICIENTE

DI CORRELAZIONE :

$$\rho_{X,Y} = \frac{\text{Cov}(X, Y)}{\sqrt{\text{Var}(X)} \sqrt{\text{Var}(Y)}}$$

OSS.

1) $x_i - \mathbb{E}[X] =$ scarto dalla media di x_i
 $\text{Cov}(X, Y) =$ somma del prodotto degli scarti dalla media

2) $\text{Cov}(X, X) = \text{Var}(X)$

3) $\text{Cov}(X, Y) = \text{Cov}(Y, X)$

4) $-1 \leq \rho_{X, Y} \leq 1$

Teorema 3.2 del Cap. "Stat. desc. e Teoremi limite"

Inoltre: $\rho_{X, Y} = \pm 1 \iff Y = aX + b$

$| \text{Cov}(X, Y) | \leq \sqrt{\text{Var}(X)} \sqrt{\text{Var}(Y)}$

5) La covarianza ci dice se c'è approssimativamente
DIPENDENZA LINEARE.

$\rho_{X,Y} = \pm 1$, allora c'è veramente dipendenza lineare.
 $\rho_{X,Y} = 0$, allora non è detto che X e Y siano
indipendenti

$$X \perp\!\!\!\perp Y \quad \xrightarrow{\text{~~NON~~}} \quad \text{Cov}(X, Y) = 0 \quad (\text{X e Y correlate})$$

Teorema

$$X \perp\!\!\!\perp Y \implies \text{Cov}(X, Y) = 0$$

Lemma

$$\text{Cov}(X, Y) = \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y]$$

Dim (Lemma)

$$\begin{aligned} \text{Cov}(X, Y) &= \mathbb{E}[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])] = \\ &= \mathbb{E}[XY - X\mathbb{E}[Y] - Y\mathbb{E}[X] + \mathbb{E}[X]\mathbb{E}[Y]] \\ &= \mathbb{E}[XY] - \mathbb{E}[Y]\mathbb{E}[X] - \mathbb{E}[X]\mathbb{E}[Y] + \mathbb{E}[X]\mathbb{E}[Y] \end{aligned}$$

N.B.

$$\mathbb{E}[aX + b] = a\mathbb{E}[X] + b$$

Dim (Teorema)

$$X \perp\!\!\!\perp Y \implies E[XY] = E[X]E[Y]$$

Quindi

$$\text{Cov}(X, Y) = E[XY] - E[X]E[Y] =$$

$$= 0$$

↑

$X \perp\!\!\!\perp Y$

$$Y = \underbrace{\log_2(1 + \sin X)}_Z$$

Ex. 2.3

X e Y v.a. discrete:

$$X \perp\!\!\!\perp Y, \quad X \sim B\left(\frac{1}{2}\right), \quad Y \sim B\left(\frac{1}{2}\right)$$

Siano $U = X + Y$ e $V = |X - Y|$

- Determinare congiunta e marginali di U e V .
- $\mathbb{P}(V < U) = ?$
- Calcolare $\text{Var}(U)$, $\text{Var}(V)$, $\text{Cov}(U, V)$.
- U e V sono indipendenti?

a)

$X \backslash Y$	0	1	P_X
0	$1/4$	$1/4$	$1/2$
1	$1/4$	$1/4$	$1/2$
P_Y	$1/2$	$1/2$	1

$$S_U = \{0, 1, 2\} \text{ e } S_V = \{0, 1\}$$

$U \backslash V$	0	1	P_U
0	$1/4$	0	$1/4$
1	0	$1/2$	$1/2$
2	$1/4$	0	$1/4$
P_V	$1/2$	$1/2$	1

(X, Y)	(U, V)
(0, 0)	(0, 0)
(0, 1)	(1, 1)
(1, 0)	(1, 1)
(1, 1)	(2, 0)

$$V \sim B\left(\frac{1}{2}\right)$$

$$U \sim \text{Bin}\left(2, \frac{1}{2}\right)$$

$$U = V \pmod{2}$$

$$b) \quad \mathbb{P}(V < U) = \sum_{\substack{i,j: \\ v_j < u_i}} P_{(U,V)}(u_i, v_j)$$

$$= P_{(U,V)}(2,0) + P_{(U,V)}(2,1) = P_U(2) = \frac{1}{4}$$

$$c) \quad \mathbb{E}[U] = 2 \cdot \frac{1}{2} = 1$$

$$\mathbb{E}[V] = \frac{1}{2}$$

$$\text{Var}(U) = 2 \cdot \frac{1}{4} = \frac{1}{2}$$

$$\text{Var}(V) = \frac{1}{4}$$

$$\mathbb{E}[UV] = \sum_{i,j} u_i v_j P_{(U,V)}(u_i, v_j) = 1 \cdot 1 \cdot P_{(U,V)}(1,1) = \frac{1}{2}$$

$$\text{Cov}(U, V) = \mathbb{E}[UV] - \mathbb{E}[U]\mathbb{E}[V] = \frac{1}{2} - \frac{1}{2} = 0$$

d) $U \perp V$?

$$U \perp V \Leftrightarrow P_{(U,V)}(u_i, v_j) = p_U(u_i) p_V(v_j)$$

No, infatti ad esempio

$$P_{(U,V)}(0,0) = \frac{1}{4} \neq \begin{matrix} P_U(0) & P_V(0) \\ \uparrow & \uparrow \\ \frac{1}{4} & \frac{1}{2} \end{matrix}$$